

Compound p -Values for Multiple Testing Procedures

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Prostate Cancer Data: Efron 2009

- $N=102$ microarrays and $M=6033$ gene expression measurements from each

m	control group				cancer group			
	1	2	...	50	1	2	...	52
1	-.931	-.840	...	3.81	-1.12	1.01	...	-.001
2	-1.07	-.880	...	-.477	-.571	-.811	...	-.836
⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮	⋮
6033	-.754	-.708	...	-.011	.457	.578	...	-.162

- **Which genes are related to cancer (differentially expressed)?**
- Need to test 6033 null hypotheses!

DATA



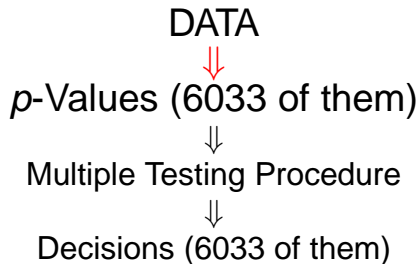
p -Values (6033 of them)



Multiple Testing Procedure



Decisions (6033 of them)



Multiple testing procedures **valid** if p -values from **null hypotheses** are

- 1 **uniformly distributed** and
- 2 **independent**

How to get a p -value for testing H_0 with data X

In words:

The p -value is “the smallest **size** allowing for the null to be rejected with the observed **data** x ”

Mathematically:

- 1 Define decision function $\delta(\mathbf{x}; \eta) \in \{0, 1\}$
 - η is size and X is data
- 2 Definition: p -Value is $P(x) = \inf\{\eta : \delta(\mathbf{x}; \eta) = 1\}$

We can work with $\delta(X; \eta)$ and use the previous definition to get a $P(X)$.

See Habiger and Pena (2011), JNS or Pena, Habiger, Wu (2011), AoS for details

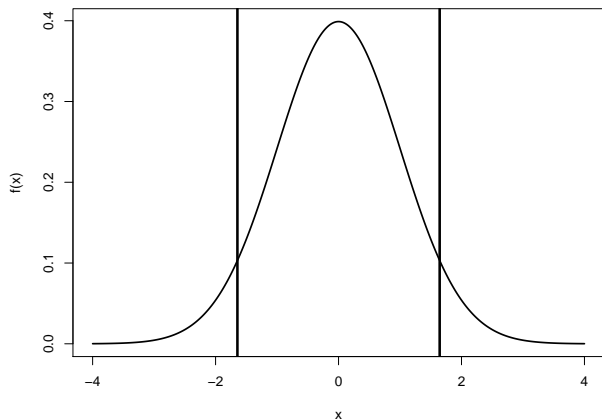
$$\begin{bmatrix} X_{11} & X_{12} & \dots & X_{1N} \\ X_{21} & X_{22} & \dots & X_{2N} \\ \vdots & \vdots & \ddots & \vdots \\ X_{M1} & X_{M2} & \dots & X_{MN} \end{bmatrix} \rightarrow \begin{bmatrix} X_1 \\ X_2 \\ \vdots \\ X_M \end{bmatrix}$$

- \rightarrow means test stat applied to each row and transformed to get $H_{0m} : X_m \sim N(0, 1)$
 - Ex. Compute $T_m = T(X_{m1}, X_{m2}, \dots, X_{mN})$ and take $X_m = \Phi^{-1}(F_T(T_m))$

- $H_{0m} : X_m \sim N(0, 1)$ vs. $H_{1m} : X_m \sim N(\mu_m, 1), \mu_m \neq 0$

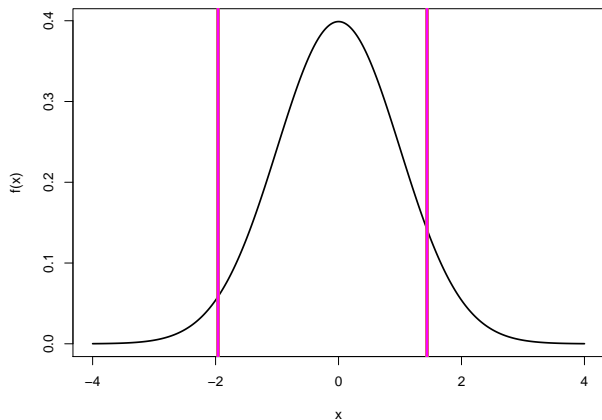
$$\delta_m(\mathbf{x}_m; \eta_m) = \begin{cases} 1 & \text{if } x_m > \Phi^{-1}(1 - \eta_m/2) \\ 1 & \text{if } x_m < \Phi^{-1}(\eta_m/2) \\ 0 & \text{otherwise} \end{cases}$$

Example



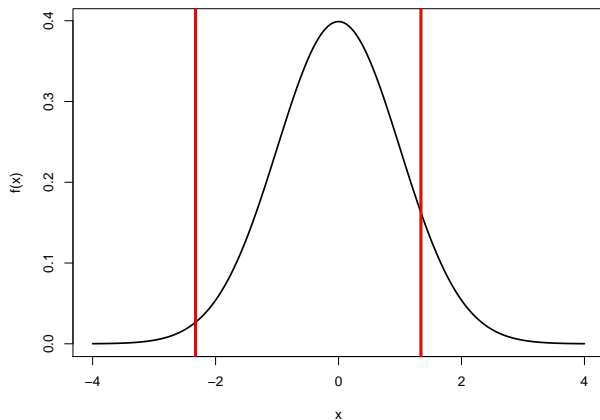
- Upper cutoff: $\Phi^{-1}(1 - .1/2) = 1.645$
- Lower Cutoff: $\Phi^{-1}(.1/2) = -1.645$
- Tail area is $\eta_m = .1!$

Example



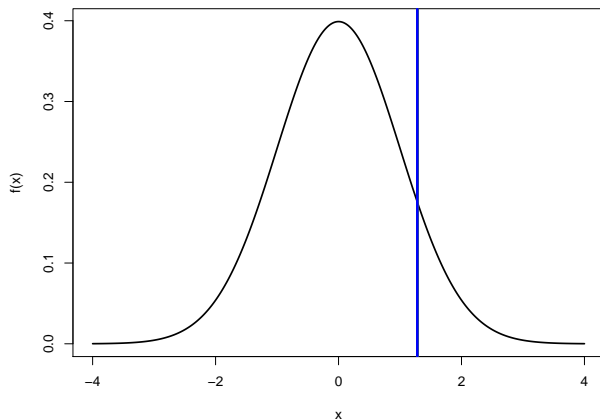
- Upper cutoff: $\Phi^{-1}(1 - .1(.75)) = 1.44$
- Lower Cutoff: $\Phi^{-1}(.1(.25)) = -1.96$
- Tail area is $\eta_m = .1!$

Example



- Upper cutoff: $\Phi^{-1}(1 - .1(.9)) = 1.34$
- Lower Cutoff: $\Phi^{-1}(.1(.1)) = -2.32$
- Tail area is $\eta_m = .1!$

Example



- Upper cutoff: $\Phi^{-1}(1 - .1(1)) = 1.28$
- Lower Cutoff: $\Phi^{-1}(.1(0)) = -\infty$
- Tail area is $\eta_m = .1!$

- **For any h_m in $[0,1]$**

$$\delta_m(\mathbf{x}_m; \eta_m) = \begin{cases} 1 & \text{if } \mathbf{x}_m > \Phi^{-1}(1 - \eta_m[1 - h_m]) \\ 1 & \text{if } \mathbf{x}_m < \Phi^{-1}(\eta_m h_m) \\ 0 & \text{otherwise} \end{cases}$$

- Optimal: $h_m(\mu_m) = I(\mu_m < 0)$
- **BUT WE DON'T KNOW $\mu_m!!!$**

$$\begin{bmatrix} X_{11} & X_{12} & \dots & X_{1N} \\ X_{21} & X_{22} & \dots & X_{2N} \\ \vdots & \vdots & \ddots & \vdots \\ X_{M1} & X_{M2} & \dots & X_{MN} \end{bmatrix} \rightarrow \begin{bmatrix} Y_1 & Z_1 \\ Y_2 & Z_2 \\ \vdots & \vdots \\ Y_M & Z_M \end{bmatrix}$$

- $\delta_1(Y_1, Y_2, \dots, Y_M, Z_1; \eta_1)$
- Plug $h_1(Y_1, Y_2, \dots, Y_M) = I(\widehat{\mu_1} < 0)$
- Z_1 is test data (rather than X_1)

$$\begin{bmatrix} X_{11} & X_{12} & \dots & X_{1N} \\ X_{21} & X_{22} & \dots & X_{2N} \\ \vdots & \vdots & \ddots & \vdots \\ X_{M1} & X_{M2} & \dots & X_{MN} \end{bmatrix} \rightarrow \begin{bmatrix} Y_1 & Z_1 \\ Y_2 & Z_2 \\ \vdots & \vdots \\ Y_M & Z_M \end{bmatrix}$$

- $\delta_2(Y_1, Y_2, \dots, Y_M, Z_2; \eta_2)$
- Plug $h_2(Y_1, Y_2, \dots, Y_M) = I(\widehat{\mu_2} < 0)$
- Z_2 is test data (rather than X_2)

$$\begin{bmatrix} X_{11} & X_{12} & \dots & X_{1N} \\ X_{21} & X_{22} & \dots & X_{2N} \\ \vdots & \vdots & \ddots & \vdots \\ X_{M1} & X_{M2} & \dots & X_{MN} \end{bmatrix} \rightarrow \begin{bmatrix} Y_1 & Z_1 \\ Y_2 & Z_2 \\ \vdots & \vdots \\ Y_M & Z_M \end{bmatrix}$$

- $\delta_M(Y_1, Y_2, \dots, Y_M, Z_M; \eta_1)$
- Plug $h_M(Y_1, Y_2, \dots, Y_M) = I(\widehat{\mu_M} < 0)$
- Z_M is test data (rather than X_M)

- The compound p -value:

- Estimate $I(\mu_m < 0)$ via $h_m(\mathbf{Y})$, where $\mathbf{Y} = (Y_1, Y_2, \dots, Y_M)$

$$P(\mathbf{Y}, Z_m) = \min \left\{ \frac{\Phi(Z_m)}{h_m(\mathbf{Y})}, \frac{1 - \Phi(Z_m)}{1 - h_m(\mathbf{Y})} \right\}$$

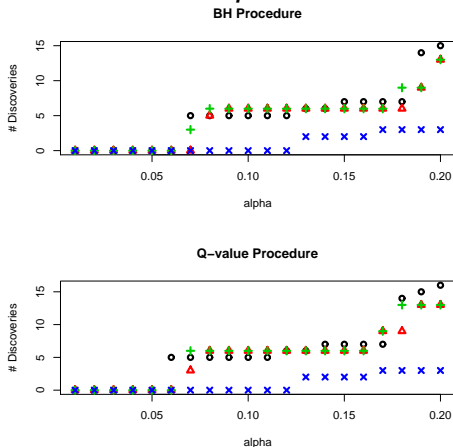
- Simple p -value:

$$P(X_m) = 2[1 - \Phi(|X_m|)] = \left\{ \frac{\Phi(X_m)}{1/2}, \frac{1 - \Phi(X_m)}{1 - 1/2} \right\}$$

- The **simple** p -value uses all the data ($X_m = Y_m + Z_m$) as test data but uses $h_m = 1/2$

Standard vs. Compound p -values

Applied Q-value and BH FDR-controlling procedures to different p -values.



- Simple: x - take $h_m = .5$, use all data as test data
- Compound: $+$, o , Δ - use 4 microarrays to estimate h_m

Theorem (simplified): If **test data** are **independent and “correctly modeled”** under nulls, these **compound p -values** are **independent and uniformly distributed** under nulls.

- “Correctly Modeled” - use T -test and $\Phi^{-1}(F_T(T_m))$ if data are normal to get Z_m s, use nonparametric test otherwise. See Habiger and Pena (2011), JNS.
- Implies many multiple testing procedures will be valid

We have ***borrowed information across tests*** to refine our upper/lower tailed cutoffs (via $h_m(\mathbf{Y})$)

- Gain in power related to “similarity” of data under H_{1mS}
 - *Theorem: If $\mu_1 = \mu_2 = \dots$ under H_{1mS} , $h_m(\mathbf{Y}) \rightarrow I(\mu_m < 0)$ as $M \rightarrow \infty$*
 - As μ_m 's become dispersed, less power is gained.

Other Approaches: Efron (2001), JASA; Sun and Cai (2007), JASA, etc.

- 1 We allow each δ_m to use different cutoffs depending on $f_1(z|\mu_m)$ under H_{m1}**
 - vs. using same cutoffs based on $f_1(z) = \int f_1(z|\mu)\pi(\mu)d\mu$
 - Is this better?
- 2 We split sample. Others double dip.**
 - efficiency vs. validity
- 3 We utilize the p -value**